

Antrittsvorlesung

Random combinatorial structures: graphs, permutations and representations

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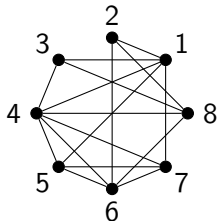


Universität
Zürich ^{UZH}

First example of random combinatorial structures: random graphs

Erdős-Rényi model of random graphs $G(n, p)$:

- G has n vertices labelled $1, \dots, n$;
- each edge $\{i, j\}$ is taken independently with probability p ;

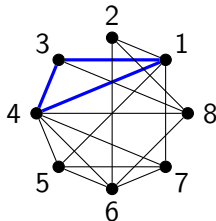


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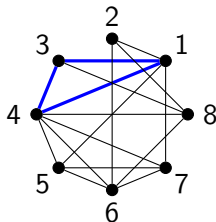
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Does G contains a **triangle**? If yes, how many?

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Does G contains a **triangle**? If yes, how many?

→ we look for an **asymptotic** answer.

Second example of random combinatorial structure: random permutation

A uniform random permutation of size 20

[9, 10, 5, 19, 7, 16, 18, 2, 14, 20, 17, 1, 6, 12, 8, 15, 11, 13, 4, 3]

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Question

Does G contains an adjacency. If yes, how many?

adjacency = consecutive values in consecutive places.

Some general motivations

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- Biological observation: from one person to another, *order of the genes* on a chromosome are the same.
From one species to another: this order changes, one can encode that by a *permutation*.
Number of adjacencies measures how close the two species are.
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→ One has to **compare with a random** permutation !
- **Spectrum of random (GOE) matrices** describe accurately the **spectra of heavy atoms** (Wigner, 50').
- **Graphs** can represent **networks**, e.g. internet (vertices are web pages, edges hyperlinks).
But Erdős-Rényi random graph is **not a good model** for internet (not the good degree distribution!).
→ A lot of **other models of random graphs** have been introduced and studied.

Outline of the talk

- 1 Presentation of the moment method
- 2 Two facets of my work related to moment methods
 - Random permutations and small cumulants
 - Random representations

First moment: description of the method

Lemma

Let X be a random variable with non-negative integer values (for example X is counting something). Then

$$P(X = 0) > 1 - E(X).$$

Proof: $E(X) = \sum_k kP(X = k) \geq \sum_{k \geq 1} P(X = k) = 1 - P(X = 0)$. \square

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Corollary (first moment method)

If $E(X_n)$ tends to 0, then $X_n = 0$ with high probability (that is, $P(X_n = 0)$ tends to 1).

First moment: application to triangles in random graphs

Let T_n be the number of triangles in $G(n, p_n)$. Then

$$T_n = \sum_{\{i,j,k\} \subset [n]} \Delta_{\{i,j,k\}},$$

where

$$\Delta_{\{i,j,k\}} = \begin{cases} 1 & \text{if } G \text{ contains the triangle } i, j, k; \\ 0 & \text{otherwise.} \end{cases}$$

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But $E(\Delta_{\{i,j,k\}}) = p_n^3$ for all $\{i, j, k\}$ and thus

$$E(T_n) = \sum_{\{i,j,k\}} E(\Delta_{\{i,j,k\}}) = \binom{n}{3} p_n^3.$$

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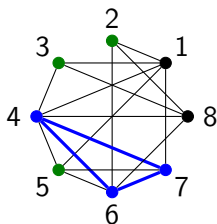
Theorem

If $n p_n \rightarrow 0$, then $G(n, p_n)$ has no triangles with high probability.

Variant: probabilistic method applied to Ramsey number.

Fix $p = 1/2$ and $k \geq 1$. Then consider X_n the number of sets W of k vertices in $G(n, 1/2)$ which are:

- either a **clique**, all pairs of vertices of W is linked by an edge;
- or an **independent set**, *i.e.* there is no edge between two vertices of W



Example : $n = 8, k = 3$

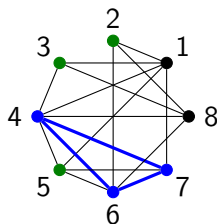
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Then

$$P(X_n = 0) > 1 - E(X_n) = 1 - 2^{1-\binom{n}{k}}$$

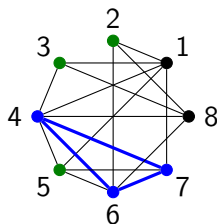


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Then

$$P(X_n = 0) > 1 - E(X_n) = 1 - 2^{1 - \binom{k}{2}} \binom{n}{k}$$

Theorem (Erdős, 1947)

If $1 - 2^{1 - \binom{k}{2}} \binom{n}{k} > 0$, then **there exists** a graph G with n vertices and neither cliques nor independent sets of size k .

Second moment: description of the method

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Lemma (Chebyshev's inequality)

$$P(|X - E(X)| \geq \lambda \sqrt{\text{Var}(X)}) \leq 1/\lambda^2.$$

In particular,

$$P(X = 0) \leq \left(\frac{E(X)}{\sqrt{\text{Var}(X)}} \right)^{-2}.$$

Proof: If Y is a non-negative r.v., then $P(Y \geq a) \leq E(Y)/a$.

Apply this to $Y = (X - E(X))^2$ and $a = \lambda^2 \text{Var}(X) = \lambda^2 E(Y)$. □

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Second moment method

If $\sqrt{\text{Var}(X_n)}/E(X_n) \rightarrow 0$, then $X_n > 0$ with high probability.

Second moment: application to triangles in random graphs

Recall that $E(T_n) \sim cn^3 p_n^3$. But

$$\begin{aligned} \text{Var}(T_n) &= E[(T_n - E(T_n))^2] = E \left[\left(\sum_{\{i,j,k\}} \Delta_{\{i,j,k\}} - E(\Delta_{\{i,j,k\}}) \right)^2 \right] \\ &= \sum_{\substack{\{i,j,k\} \\ \{i',j',k'\}}} E [(\Delta_{\{i,j,k\}} - E(\Delta_{\{i,j,k\}})) \cdot (\Delta_{\{i',j',k'\}} - E(\Delta_{\{i',j',k'\}}))] \end{aligned}$$

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But this is true for most of the terms in the sum! One can show:

Theorem

If $np_n \rightarrow \infty$, then $\sqrt{\text{Var}(T_n)}/E(T_n)$ tends to 0 and thus $G(n, p_n)$ contains a triangle with high probability.

Transition

- To prove that our random structure **does not contain** a given type of substructure:

set X_n the **number** of these substructures $X_n = \sum_{\alpha} I_{\alpha}$.

Compute $E(X_n)$ by **linearity** and show that it tends to 0.

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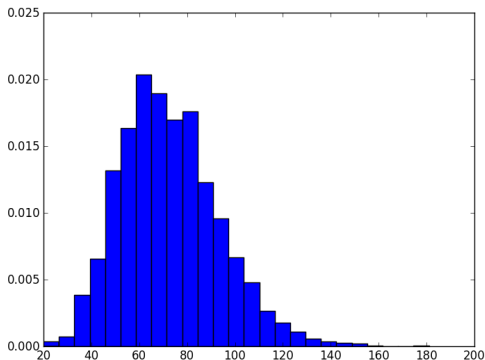
$$\lim_{n \rightarrow \infty} \sqrt{\text{Var}(T_n)} / E(T_n) = 0.$$

Question

In the case where there is a triangle with high probability, **how many** are there ?

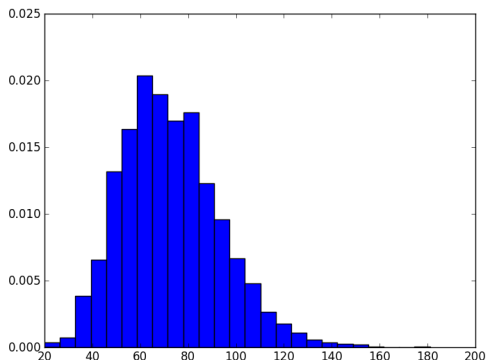
Number of triangles in random graphs: a simulation

Here is a histogram number of triangles in 5000 independent random graphs $G(20, .4)$.



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Theorem (Ruciński, 1988)

If $np_n \rightarrow \infty$ and $n^2(1 - p_n) \rightarrow \infty$, then T_n is asymptotically Gaussian.

The moment method

Theorem (the moment method)

Let X_n be a sequence of random variable and N a standard normal random variable. If for each $\ell \geq 1$,

$$\lim_{n \rightarrow \infty} E(X_n^\ell) = E(N^\ell),$$

then X_n converges to N in distribution.

Remark:

$$E(N^\ell) = \begin{cases} 1 \cdot 3 \cdots (2m - 1) & \text{if } \ell = 2m \text{ is even;} \\ 0 & \text{if } \ell \text{ is odd.} \end{cases}$$

Sketch of proof of Ruciński theorem

Let T_n be the number of triangles in $G(n, p_n)$ and $X_n = \frac{T_n - E(T_n)}{\sqrt{\text{Var}(T_n)}}$. Then

$$E(X_n^\ell) = \frac{1}{\text{Var}(T_n)^{\ell/2}} E\left((T_n - E(T_n))^\ell\right)$$

But $T_n - E(T_n) = \sum_{\{i,j,k\}} \Delta_{\{i,j,k\}} - E(\Delta_{\{i,j,k\}})$. Thus

$$\text{Var}(T_n)^{\ell/2} E(X_n^\ell) = \sum_{\substack{\{i_1, j_1, k_1\} \\ \vdots \\ \{i_\ell, j_\ell, k_\ell\}}} E\left[\left(\Delta_{\{i_1, j_1, k_1\}} - E(\Delta_{\{i_1, j_1, k_1\}})\right) \cdots \left(\Delta_{\{i_\ell, j_\ell, k_\ell\}} - E(\Delta_{\{i_\ell, j_\ell, k_\ell\}})\right)\right]$$

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Because of independence, most of the terms vanish... One has to study asymptotics carefully and find $\lim E(X_n^\ell) = E(N^\ell)$.

A TCL for the number of prime divisors

Each integer $X \geq 1$ writes uniquely as a product of primes:

$$18984 = 2^3 * 3 * 7 * 113.$$

Theorem (Erdős, Kac, 1940)

Let D_n be the number of prime divisors of a uniform random integer between 1 and n . Then

$$X_n := \frac{D_n - E(D_n)}{\sqrt{\text{Var}(D_n)}} \xrightarrow{d} N.$$

$$D_n = \sum_{\substack{p \text{ prime} \\ p \leq n}} I_p(n) \text{ where } I_p(n) = \begin{cases} 1 & \text{if } p \text{ divides } n; \\ 0 & \text{otherwise.} \end{cases}$$

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No independence, but, for distinct primes,

$$E(I_{p_1} \dots I_{p_k}) = \frac{1}{n} \left\lfloor \frac{n}{p_1 \dots p_k} \right\rfloor \simeq \frac{1}{\underbrace{p_1 \dots p_k}} .$$

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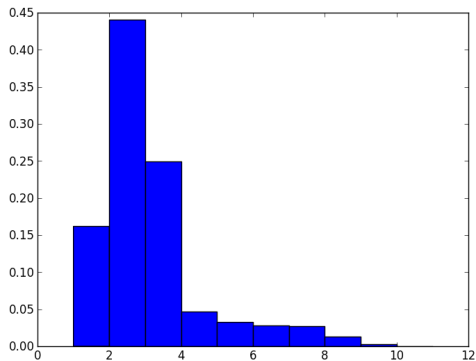
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$\Rightarrow D_n$ also converges after normalization to a normal law.

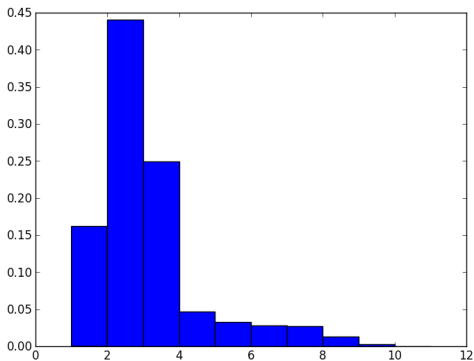
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$\ln(\ln(n)) \sim 5$, but the empirical mean is 2.5.

Erdős-Kac theorem is an **asymptotic** result!

Random permutations and non independence

Definition

A permutation of n is a word with letters from 1 to n , which contains exactly one each letter.

A uniform random permutation of size 20:

[9, 10, 5, 19, 7, 16, 18, 2, 14, 20, 17, 1, 6, 12, 8, 15, 11, 13, 4, 3]

The first element σ_1 is uniform between 1 and n . Also the second σ_2, \dots
But they are not independent.

Measuring their dependence

Set $\sigma'_1 = \sigma_1/n$, $\sigma'_2 = \sigma_2/n$. Their *covariance* is

$$\text{Cov}(\sigma'_1, \sigma'_2) := E(\sigma'_1 \sigma'_2) - E(\sigma'_1)E(\sigma'_2) = \dots = O(1/n).$$

Small correlation!

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$$\begin{aligned} \kappa_3(\sigma'_1, \sigma'_2, \sigma'_3) &= E(\sigma'_1 \sigma'_2 \sigma'_3) - E(\sigma'_1 \sigma'_2)E(\sigma'_3) - \dots + 2E(\sigma'_1)E(\sigma'_2)E(\sigma'_3) \\ &= \dots = O(1/n^2). \end{aligned}$$

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For “local” statistics, allows to use moment method *as if* $\sigma_1, \sigma_2, \dots$ were *independent*.

Example: the number of adjacencies is asymptotically Poisson distributed!

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For “local” statistics, allows to use moment method *as if* $\sigma_1, \sigma_2, \dots$ were independent.

Small cumulants also appear in random graphs with fixed number of edges, random orthogonal/unitary matrices, ...

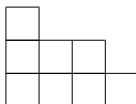
Partitions

Definition

A **partition** (of n) is a non-increasing list of integer (of sum n).

Example : $(4, 3, 1)$ is a partition of 8.

Representation as **Young diagram** :



Plancherel measure

Representation theory of symmetric group associates to each partition λ of n a vector space V_λ , called **representation**, such that

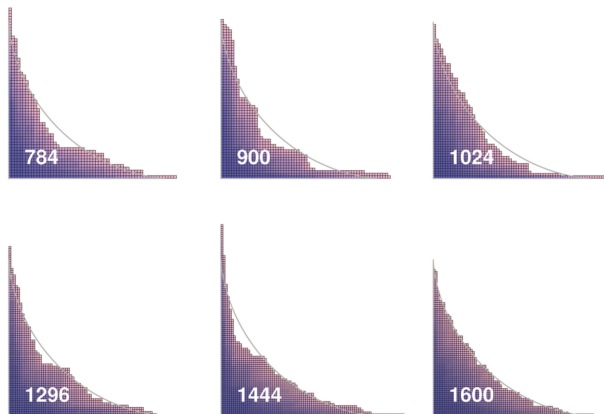
$$\sum_{\substack{\lambda \text{ partition} \\ \text{of } n}} (\dim V_\lambda)^2 = n!$$

We consider Plancherel measure on partitions of size n :

$$P(\lambda) = \frac{(\dim V_\lambda)^2}{n!}.$$

This defines a model for a random partition of size n , but **not uniform**.

Random partitions under Plancherel measure: simulation



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Nice limit shape!

Moment method on characters

Lemma (easy algebraic statement rewritten in probabilistic terms)

$$E(\hat{\chi}^\bullet(\sigma)) = \begin{cases} 1 & \text{if } \sigma = \text{id}; \\ 0 & \text{otherwise.} \end{cases},$$

where $\hat{\chi}$ are **irreducible characters**.

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Algebraic combinatorics provides a lot of non-trivial probabilistic models that can be analysed by moment method \rightarrow **integrable probability**.