

(Weighted) dependency graphs and asymptotic normality

Valentin Féray

CNRS, Institut Élie Cartan de Lorraine (IECL)

L^2 workshop in Probability and Statistics
Luxembourg, September 15th, 2022



What is this talk about ?

Consider some sequence of r.v. X_n (e.g., number of substructures of a given type in some probabilistic model).

Goal: prove that some X_n satisfies is **asymptotically normal**, i.e.

$$\frac{X_n - \mathbb{E}[X_n]}{\sqrt{\text{Var}(X_n)}} \xrightarrow{d} \mathcal{N}(0, 1).$$

What is this talk about ?

Consider some sequence of r.v. X_n (e.g., number of substructures of a given type in some probabilistic model).

Goal: prove that some X_n satisfies is **asymptotically normal**, i.e.

$$\frac{X_n - \mathbb{E}[X_n]}{\sqrt{\text{Var}(X_n)}} \xrightarrow{d} \mathcal{N}(0, 1).$$

Available tools:

- computation of characteristic functions;
- Stein's method;
- **moment (or cumulant) methods**

What is this talk about ?

Consider some sequence of r.v. X_n (e.g., number of substructures of a given type in some probabilistic model).

Goal: prove that some X_n satisfies is **asymptotically normal**, i.e.

$$\frac{X_n - \mathbb{E}[X_n]}{\sqrt{\text{Var}(X_n)}} \xrightarrow{d} \mathcal{N}(0, 1).$$

Available tools:

- computation of characteristic functions;
- Stein's method;
- **moment (or cumulant) methods**

Today: **(weighted) dependency graphs**, based on cumulants and independence (or weak dependencies) between variables.

Outline of the talk

- 1 Dependency graphs
 - A motivating example: substrings in random words
 - An asymptotic normality criterion
- 2 Weighted dependency graphs
 - Definition and an extended normality criterion
 - Back to subwords: Markovian texts

Substrings in random words (1/2)

(following Flajolet, Guivarc'h, Szpankowski, and Vallée, '01)

Let w be a **random word** of size n with **independent** (identically distributed) letters taken in a finite alphabet \mathcal{A} .

Fix a word u , called "pattern" of length ℓ .

An **occurrence** of u in w is a ℓ -tuple $i_1 < \dots < i_\ell$ s.t. $w_{i_1} = u_1, \dots, w_{i_\ell} = u_\ell$.

Example: two occurrences of aab in $w = \underline{a}ab\underline{b}ba\underline{a}ab$ (one in blue, one underlined)

(Variants: consecutive occurrences, allowing gaps of given lengths).

Substrings in random words (1/2)

(following Flajolet, Guivarc'h, Szpankowski, and Vallée, '01)

Let w be a **random word** of size n with **independent** (identically distributed) letters taken in a finite alphabet \mathcal{A} .

Fix a word u , called "pattern" of length ℓ .

An **occurrence** of u in w is a ℓ -tuple $i_1 < \dots < i_\ell$ s.t. $w_{i_1} = u_1, \dots, w_{i_\ell} = u_\ell$.

Example: two occurrences of aab in $w = a\underline{a}b\underline{b}a\underline{a}b$ (one in blue, one underlined)

Question

Asymptotic behaviour of the number X_n of occurrences of u in w ?

Motivations: intrusion detection in computer science, discovering meaningful strings of DNA, ...

Substrings in random words (2/2)

Theorem (FGSV, '01)

We have

$$\mathbb{E}[X_n] \sim C_1 n^\ell, \quad \text{Var}[X_n] = (C_2 + o(1))n^{2\ell-1},$$

where $C_1 > 0$ and C_2 are computable constants.

Moreover, if $C_2 > 0$, then X_n is asymptotically normal.

Substrings in random words (2/2)

Theorem (FGSV, '01)

We have

$$\mathbb{E}[X_n] \sim C_1 n^\ell, \quad \text{Var}[X_n] = (C_2 + o(1))n^{2\ell-1},$$

where $C_1 > 0$ and C_2 are computable constants.

Moreover, if $C_2 > 0$, then X_n is asymptotically normal.

The proof of the asymptotic normality uses the method of moments.

I will sketch it using [cumulants and dependency graphs](#) (essentially the same proof, but presented differently, and in a general context).

Substrings in random words (2/2)

Theorem (FGSV, '01)

We have

$$\mathbb{E}[X_n] \sim C_1 n^\ell, \quad \text{Var}[X_n] = (C_2 + o(1))n^{2\ell-1},$$

where $C_1 > 0$ and C_2 are computable constants.

Moreover, if $C_2 > 0$, then X_n is asymptotically normal.

The proof of the asymptotic normality uses the method of moments.

I will sketch it using [cumulants and dependency graphs](#) (essentially the same proof, but presented differently, and in a general context).

Notation: for $I \subseteq [n]$, $|I| = \ell$, set $Y_I = \mathbf{1}[u \text{ occurs at position } I \text{ in } \mathbf{w}]$.

Then $X_n = \sum_{I \in \binom{[n]}{\ell}} Y_I$.

Transition

1 Dependency graphs

- A motivating example: substrings in random words
- An asymptotic normality criterion

2 Weighted dependency graphs

- Definition and an extended normality criterion
- Back to subwords: Markovian texts

Dependency graphs

Definition (Malyshev, '80, Petrovskaya/Leontovich, '82, Janson, '88)

A graph L with vertex set A is a dependency graph for the family $\{Y_\alpha, \alpha \in A\}$ if the following holds for any $A_1, A_2 \subset A$:

there is no edge
between A_1 and A_2 \implies $\{Y_\alpha, \alpha \in A_1\}$ and $\{Y_\alpha, \alpha \in A_2\}$
are independent

Roughly: there is an edge between pairs of **dependent** random variables.

Dependency graphs

Definition (Malyshev, '80, Petrovskaya/Leontovich, '82, Janson, '88)

A graph L with vertex set A is a dependency graph for the family $\{Y_\alpha, \alpha \in A\}$ if the following holds for any $A_1, A_2 \subset A$:

there is no edge \implies $\{Y_\alpha, \alpha \in A_1\}$ and $\{Y_\alpha, \alpha \in A_2\}$
 between A_1 and A_2 are independent

Roughly: there is an edge between pairs of **dependent** random variables.

Example

Consider our random word problem. Let $A = \binom{[n]}{\ell}$ and

$$\{I_1, I_2\} \in E_L \text{ iff } I_1 \cap I_2 \neq \emptyset.$$

Then L is a **dependency graph** for the family $\{Y_I, I \in \binom{[n]}{\ell}\}$.

Dependency graphs

Definition (Malyshev, '80, Petrovskaya/Leontovich, '82, Janson, '88)

A graph L with vertex set A is a dependency graph for the family $\{Y_\alpha, \alpha \in A\}$ if the following holds for any $A_1, A_2 \subset A$:

there is no edge
between A_1 and A_2 \implies $\{Y_\alpha, \alpha \in A_1\}$ and $\{Y_\alpha, \alpha \in A_2\}$
are independent

Roughly: there is an edge between pairs of **dependent** random variables.

Example

Note: L is regular of degree $\mathcal{O}(n^{\ell-1})$

Consider our random word problem. Let $A = \binom{[n]}{\ell}$ and

$\{I_1, I_2\} \in E_L$ iff $I_1 \cap I_2 \neq \emptyset$.

Then L is a **dependency graph** for the family $\{Y_I, I \in \binom{[n]}{\ell}\}$.

Janson's normality criterion

Setting: for each n ,

- $\{Y_{n,i}, 1 \leq i \leq N_n\}$ is a family of bounded random variables; $|Y_{n,i}| < M$ a.s.
- we have a dependency graph L_n with maximal degree $D_n - 1$.
- we set $X_n = \sum_{i=1}^{N_n} Y_{n,i}$ and $\sigma_n^2 = \text{Var}(X_n)$.

Janson's normality criterion

Setting: for each n ,

- $\{Y_{n,i}, 1 \leq i \leq N_n\}$ is a family of bounded random variables; $|Y_{n,i}| < M$ a.s.
- we have a dependency graph L_n with maximal degree $D_n - 1$.
- we set $X_n = \sum_{i=1}^{N_n} Y_{n,i}$ and $\sigma_n^2 = \text{Var}(X_n)$.

Theorem (Janson, 1988)

Assume that $\left(\frac{N_n}{D_n}\right)^{1/s} \frac{D_n}{\sigma_n} \rightarrow 0$ for some integer s .

Then X_n is asymptotically normal.

Janson's normality criterion

Setting: for each n ,

- $\{Y_{n,i}, 1 \leq i \leq N_n\}$ is a family of bounded random variables; $|Y_{n,i}| < M$ a.s.
- we have a dependency graph L_n with maximal degree $D_n - 1$.
- we set $X_n = \sum_{i=1}^{N_n} Y_{n,i}$ and $\sigma_n^2 = \text{Var}(X_n)$.

Theorem (Janson, 1988)

Assume that $\left(\frac{N_n}{D_n}\right)^{1/s} \frac{D_n}{\sigma_n} \rightarrow 0$ for some integer s .

Then X_n is asymptotically normal.

Example: For occurrences of u in \mathbf{w} , we have

$$N_n = \Theta(n^\ell), D_n = \Theta(n^{\ell-1}) \text{ and } \sigma_n = \Theta(n^{\ell-1/2}),$$

so that asymptotic normality follows (assuming the variance estimates!).

Janson's normality criterion

Setting: for each n ,

- $\{Y_{n,i}, 1 \leq i \leq N_n\}$ is a family of bounded random variables; $|Y_{n,i}| < M$ a.s.
- we have a **dependency graph** L_n with maximal degree $D_n - 1$.
- we set $X_n = \sum_{i=1}^{N_n} Y_{n,i}$ and $\sigma_n^2 = \text{Var}(X_n)$.

Theorem (Janson, 1988)

Assume that $\left(\frac{N_n}{D_n}\right)^{1/s} \frac{D_n}{\sigma_n} \rightarrow 0$ for some integer s .

Then X_n is asymptotically normal.

In roughly the same setting, we also have **bounds on the speed of convergence** and **deviation estimates** (see Baldi, Rinott, '89, Rinott, '94 and F., Méliot, Nikeghbali, '16, '17).

Main tool in the proof: (mixed) cumulants

- **Definition:** mixed cumulants are multilinear functionals defined by

$$\kappa_r(X_1, \dots, X_r) = [t_1 \cdots t_r] \log \left(\mathbb{E} \left[\exp \left(\sum_{j=1}^r t_j X_j \right) \right] \right).$$

Examples:

$$\begin{aligned} \kappa_1(X) &:= \mathbb{E}(X), & \kappa_2(X, Y) &:= \text{Cov}(X, Y) = \mathbb{E}(XY) - \mathbb{E}(X)\mathbb{E}(Y) \\ \kappa_3(X, Y, Z) &:= \mathbb{E}(XYZ) - \mathbb{E}(XY)\mathbb{E}(Z) - \mathbb{E}(XZ)\mathbb{E}(Y) \\ &\quad - \mathbb{E}(YZ)\mathbb{E}(X) + 2\mathbb{E}(X)\mathbb{E}(Y)\mathbb{E}(Z). \end{aligned}$$

Note: $\kappa_\ell(X) := \kappa_\ell(X, \dots, X)$ is the usual cumulant of a single r.v.

Main tool in the proof: (mixed) cumulants

- **Definition:** mixed cumulants are multilinear functionals defined by

$$\kappa_r(X_1, \dots, X_r) = [t_1 \cdots t_r] \log \left(\mathbb{E} \left[\exp \left(\sum_{j=1}^r t_j X_j \right) \right] \right).$$

Examples:

$$\begin{aligned} \kappa_1(X) &:= \mathbb{E}(X), & \kappa_2(X, Y) &:= \text{Cov}(X, Y) = \mathbb{E}(XY) - \mathbb{E}(X)\mathbb{E}(Y) \\ \kappa_3(X, Y, Z) &:= \mathbb{E}(XYZ) - \mathbb{E}(XY)\mathbb{E}(Z) - \mathbb{E}(XZ)\mathbb{E}(Y) \\ &\quad - \mathbb{E}(YZ)\mathbb{E}(X) + 2\mathbb{E}(X)\mathbb{E}(Y)\mathbb{E}(Z). \end{aligned}$$

Note: $\kappa_\ell(X) := \kappa_\ell(X, \dots, X)$ is the usual cumulant of a single r.v.

- If a set of variables can be split in two mutually independent sets, then its mixed cumulant vanishes.

Main tool in the proof: (mixed) cumulants

- **Definition:** mixed cumulants are multilinear functionals defined by

$$\kappa_r(X_1, \dots, X_r) = [t_1 \cdots t_r] \log \left(\mathbb{E} \left[\exp \left(\sum_{j=1}^r t_j X_j \right) \right] \right).$$

Examples:

$$\begin{aligned} \kappa_1(X) &:= \mathbb{E}(X), & \kappa_2(X, Y) &:= \text{Cov}(X, Y) = \mathbb{E}(XY) - \mathbb{E}(X)\mathbb{E}(Y) \\ \kappa_3(X, Y, Z) &:= \mathbb{E}(XYZ) - \mathbb{E}(XY)\mathbb{E}(Z) - \mathbb{E}(XZ)\mathbb{E}(Y) \\ &\quad - \mathbb{E}(YZ)\mathbb{E}(X) + 2\mathbb{E}(X)\mathbb{E}(Y)\mathbb{E}(Z). \end{aligned}$$

Note: $\kappa_\ell(X) := \kappa_\ell(X, \dots, X)$ is the usual cumulant of a single r.v.

- If a set of variables can be split in two mutually independent sets, then its mixed cumulant vanishes.
- Let $\sigma_n = \sqrt{\text{Var}(X_n)}$. If, for some $s \geq 3$ and any $r \geq s$, we have $\kappa_r(X_n) = o(\sigma_n^r)$, then X_n is asymptotically normal. (Janson, 1988)

Sketch of proof of Janson's normality criterion

Setting: for each n ,

- $\{Y_{n,i}, 1 \leq i \leq N_n\}$ is a family of bounded r.v.; $|Y_{n,i}| < M$ a.s.
- we have a **dependency graph** L_n with maximal degree $D_n - 1$.
- we set $X_n = \sum_{i=1}^{N_n} Y_{n,i}$ and $\sigma_n^2 = \text{Var}(X_n)$.
- we assume $\left(\frac{N_n}{D_n}\right)^{1/s} \frac{D_n}{\sigma_n} \rightarrow 0$ for some $s \geq 3$.

Sketch of proof of Janson's normality criterion

Setting: for each n ,

- $\{Y_{n,i}, 1 \leq i \leq N_n\}$ is a family of bounded r.v.; $|Y_{n,i}| < M$ a.s.
- we have a **dependency graph** L_n with maximal degree $D_n - 1$.
- we set $X_n = \sum_{i=1}^{N_n} Y_{n,i}$ and $\sigma_n^2 = \text{Var}(X_n)$.
- we assume $\left(\frac{N_n}{D_n}\right)^{1/s} \frac{D_n}{\sigma_n} \rightarrow 0$ for some $s \geq 3$.

Fix $r \geq 1$. Then

$$\kappa_r(X_n) = \sum_{i_1, \dots, i_r} \kappa(Y_{n,i_1}, \dots, Y_{n,i_r}).$$

Sketch of proof of Janson's normality criterion

Setting: for each n ,

- $\{Y_{n,i}, 1 \leq i \leq N_n\}$ is a family of bounded r.v.; $|Y_{n,i}| < M$ a.s.
- we have a **dependency graph** L_n with maximal degree $D_n - 1$.
- we set $X_n = \sum_{i=1}^{N_n} Y_{n,i}$ and $\sigma_n^2 = \text{Var}(X_n)$.
- we assume $\left(\frac{N_n}{D_n}\right)^{1/s} \frac{D_n}{\sigma_n} \rightarrow 0$ for some $s \geq 3$.

Fix $r \geq 1$. Then

$$\kappa_r(X_n) = \sum_{i_1, \dots, i_r} \kappa(Y_{n,i_1}, \dots, Y_{n,i_r}).$$

Each summand is 0, unless **the induced graph** $L_n[i_1, \dots, i_r]$ is connected.

Sketch of proof of Janson's normality criterion

Setting: for each n ,

- $\{Y_{n,i}, 1 \leq i \leq N_n\}$ is a family of bounded r.v.s; $|Y_{n,i}| < M$ a.s.
- we have a **dependency graph** L_n with maximal degree $D_n - 1$.
- we set $X_n = \sum_{i=1}^{N_n} Y_{n,i}$ and $\sigma_n^2 = \text{Var}(X_n)$.
- we assume $\left(\frac{N_n}{D_n}\right)^{1/s} \frac{D_n}{\sigma_n} \rightarrow 0$ for some $s \geq 3$.

Fix $r \geq 1$. Then

$$\kappa_r(X_n) = \sum_{i_1, \dots, i_r} \kappa(Y_{n,i_1}, \dots, Y_{n,i_r}).$$

Each summand is 0, unless, **up to reordering**, each i_j is a neighbour of either i_1, \dots , or i_{j-1} .

Sketch of proof of Janson's normality criterion

Setting: for each n ,

- $\{Y_{n,i}, 1 \leq i \leq N_n\}$ is a family of bounded r.v.; $|Y_{n,i}| < M$ a.s.
- we have a dependency graph L_n with maximal degree $D_n - 1$.
- we set $X_n = \sum_{i=1}^{N_n} Y_{n,i}$ and $\sigma_n^2 = \text{Var}(X_n)$.
- we assume $\left(\frac{N_n}{D_n}\right)^{1/s} \frac{D_n}{\sigma_n} \rightarrow 0$ for some $s \geq 3$.

Fix $r \geq 1$. Then

$$\kappa_r(X_n) = \sum_{i_1, \dots, i_r} \kappa(Y_{n,i_1}, \dots, Y_{n,i_r}).$$

Each summand is 0, unless, up to reordering, each i_j is a neighbour of either i_1, \dots , or i_{j-1} . We have $r!$ choices for the reordering, N_n choices for i_1 , D_n choices for i_2 , $2D_n$ choices for i_3, \dots

→ at most $(r!)^2 N_n D_n^{r-1}$ non-zero terms, each of which is bounded by $C_r M^r$.

Sketch of proof of Janson's normality criterion

Setting: for each n ,

- $\{Y_{n,i}, 1 \leq i \leq N_n\}$ is a family of bounded r.v.; $|Y_{n,i}| < M$ a.s.
- we have a dependency graph L_n with maximal degree $D_n - 1$.
- we set $X_n = \sum_{i=1}^{N_n} Y_{n,i}$ and $\sigma_n^2 = \text{Var}(X_n)$.
- we assume $\left(\frac{N_n}{D_n}\right)^{1/s} \frac{D_n}{\sigma_n} \rightarrow 0$ for some $s \geq 3$.

Fix $r \geq 1$. Then

$$\kappa_r(X_n) = \sum_{i_1, \dots, i_r} \kappa(Y_{n,i_1}, \dots, Y_{n,i_r}).$$

→ at most $(r!)^2 N_n D_n^{r-1}$ non-zero terms, each of which is bounded by $C_r M^r$.

$$\begin{aligned} |\kappa_r(X_n)| &\leq C_r (r!)^2 N_n D_n^{r-1} M^r \\ &= o(\sigma_n^r) \quad (\text{for } r \geq s, \text{ using the assumption}) \quad \square \end{aligned}$$

Applications of dependency graphs to asymptotic normality results

- mathematical modelization of cell populations (Petrovskaya, Leontovich, '82);
- subgraph counts in random graphs (Janson, Baldi, Rinott, Penrose, '88, '89, '95, '03);
- Geometric probability: length of k neighbour graphs of random points, ... (Avram, Bertsimas, Penrose, Yukich, Bárány, Vu, '93, '05, '07);
- pattern occurrences in random permutations (Bóna, Janson, Hitchenko, Nakamura, Zeilberger, Hofer, '07, '09, '14, '18).

Transition

- 1 Dependency graphs
 - A motivating example: substrings in random words
 - An asymptotic normality criterion
- 2 Weighted dependency graphs
 - Definition and an extended normality criterion
 - Back to subwords: Markovian texts

Motivation: models with "weak dependencies"

In many models, we do not have independence, but only *weak dependencies*:

- subword occurrences in a text generated by a [Markovian source](#);

Motivation: models with "weak dependencies"

In many models, we do not have independence, but only *weak dependencies*:

- subword occurrences in a text generated by a [Markovian source](#);
- subgraph counts in uniform random graphs with [fixed number of edges](#) ($G(n, M)$) or [fixed vertex degrees](#);

Motivation: models with "weak dependencies"

In many models, we do not have independence, but only *weak dependencies*:

- subword occurrences in a text generated by a **Markovian source**;
- subgraph counts in uniform random graphs with **fixed number of edges** ($G(n, M)$) or **fixed vertex degrees**;
- "local statistics" $\sum_{i_1, \dots, i_r} F(i_1, \dots, i_r, \pi(i_1), \dots, \pi(i_r))$ in a **uniform random permutation** π ;

Motivation: models with "weak dependencies"

In many models, we do not have independence, but only *weak dependencies*:

- subword occurrences in a text generated by a **Markovian source**;
- subgraph counts in uniform random graphs with **fixed number of edges** ($G(n, M)$) or **fixed vertex degrees**;
- "local statistics" $\sum_{i_1, \dots, i_r} F(i_1, \dots, i_r, \pi(i_1), \dots, \pi(i_r))$ in a **uniform random permutation** π ;
- statistical physics: exclusion process (SSEP), **Ising model**.

Motivation: models with "weak dependencies"

In many models, we do not have independence, but only *weak dependencies*:

- subword occurrences in a text generated by a [Markovian source](#);
- subgraph counts in uniform random graphs with [fixed number of edges](#) ($G(n, M)$) or [fixed vertex degrees](#);
- "local statistics" $\sum_{i_1, \dots, i_r} F(i_1, \dots, i_r, \pi(i_1), \dots, \pi(i_r))$ in a [uniform random permutation](#) π ;
- statistical physics: exclusion process (SSEP), [Ising model](#).

Goal: [extend Janson's normality criterion](#), to cover the above frameworks.

Weighted dependency graphs

We use weighted graphs, i.e. graphs with a weight in $[0, 1]$ on each edge (weight 0 \equiv no edge).

Definition (F., '18)

Fix $\mathbf{C} = (C_r)_{r \geq 1}$. A weighted graph \tilde{L} with vertex set A is a **C-weighted dependency graph** for the family $\{Y_\alpha, \alpha \in A\}$ if, for any $\alpha_1, \dots, \alpha_r$ in A ,

$$|\kappa(Y_{\alpha_1}, \dots, Y_{\alpha_r})| \leq C_r \mathcal{M}(\tilde{L}[\alpha_1, \dots, \alpha_r]).$$

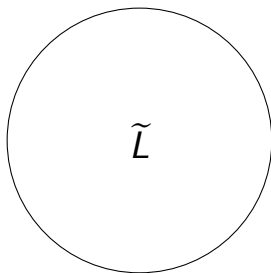
Weighted dependency graphs

We use weighted graphs, i.e. graphs with a weight in $[0, 1]$ on each edge (weight 0 \equiv no edge).

Definition (F., '18)

Fix $\mathbf{C} = (C_r)_{r \geq 1}$. A weighted graph \tilde{L} with vertex set A is a **C-weighted dependency graph** for the family $\{Y_\alpha, \alpha \in A\}$ if, for any $\alpha_1, \dots, \alpha_r$ in A ,

$$|\kappa(Y_{\alpha_1}, \dots, Y_{\alpha_r})| \leq C_r \mathcal{M}(\tilde{L}[\alpha_1, \dots, \alpha_r]).$$



Weighted dependency graphs

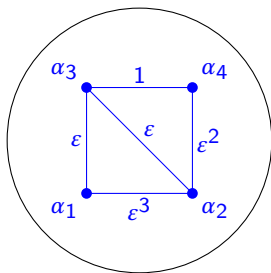
We use weighted graphs, i.e. graphs with a weight in $[0, 1]$ on each edge (weight $0 \equiv$ no edge).

Definition (F., '18)

Fix $\mathbf{C} = (C_r)_{r \geq 1}$. A weighted graph \tilde{L} with vertex set A is a **C-weighted dependency graph** for the family $\{Y_\alpha, \alpha \in A\}$ if, for any $\alpha_1, \dots, \alpha_r$ in A ,

$$|\kappa(Y_{\alpha_1}, \dots, Y_{\alpha_r})| \leq C_r \mathcal{M}(\tilde{L}[\alpha_1, \dots, \alpha_r]).$$

$\tilde{L}[\alpha_1, \dots, \alpha_r]$: graph induced by \tilde{L} on vertices $\alpha_1, \dots, \alpha_r$.



Weighted dependency graphs

We use weighted graphs, i.e. graphs with a weight in $[0, 1]$ on each edge (weight 0 \equiv no edge).

Definition (F., '18)

Fix $\mathbf{C} = (C_r)_{r \geq 1}$. A weighted graph \tilde{L} with vertex set A is a **C-weighted dependency graph** for the family $\{Y_\alpha, \alpha \in A\}$ if, for any $\alpha_1, \dots, \alpha_r$ in A ,

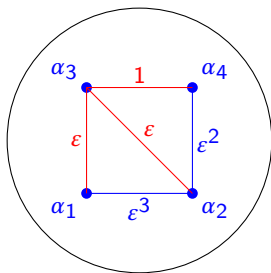
$$|\kappa(Y_{\alpha_1}, \dots, Y_{\alpha_r})| \leq C_r \mathcal{M}(\tilde{L}[\alpha_1, \dots, \alpha_r]).$$

$\tilde{L}[\alpha_1, \dots, \alpha_r]$: graph induced by \tilde{L} on vertices $\alpha_1, \dots, \alpha_r$.

$\mathcal{M}(K)$: Maximum weight of a spanning tree of K (= product of the edge weights).

In the example,

$$\mathcal{M}(\tilde{L}[\alpha_1, \dots, \alpha_4]) = \varepsilon^2.$$



Weighted dependency graphs

We use weighted graphs, i.e. graphs with a weight in $[0, 1]$ on each edge (weight 0 \equiv no edge).

Definition (F., '18)

Fix $\mathbf{C} = (C_r)_{r \geq 1}$. A weighted graph \tilde{L} with vertex set A is a **C-weighted dependency graph** for the family $\{Y_\alpha, \alpha \in A\}$ if, for any $\alpha_1, \dots, \alpha_r$ in A ,

$$|\kappa(Y_{\alpha_1}, \dots, Y_{\alpha_r})| \leq C_r \mathcal{M}(\tilde{L}[\alpha_1, \dots, \alpha_r]).$$

Intuition: the smaller the edge weights are, the smaller the cumulant should be. The **edge weights quantify the dependencies** between variables.

Weighted dependency graphs

We use weighted graphs, i.e. graphs with a weight in $[0, 1]$ on each edge (weight 0 \equiv no edge).

Definition (F., '18)

Fix $\mathbf{C} = (C_r)_{r \geq 1}$. A weighted graph \tilde{L} with vertex set A is a **C-weighted dependency graph** for the family $\{Y_\alpha, \alpha \in A\}$ if, for any $\alpha_1, \dots, \alpha_r$ in A ,

$$|\kappa(Y_{\alpha_1}, \dots, Y_{\alpha_r})| \leq C_r \mathcal{M}(\tilde{L}[\alpha_1, \dots, \alpha_r]).$$

Intuition: the smaller the edge weights are, the smaller the cumulant should be. The **edge weights quantify the dependencies** between variables.

⚠ Unlike for usual dependency graphs, **proving that something is a weighted dependency graph needs work!**

Weighted dependency graphs

We use weighted graphs, i.e. graphs with a weight in $[0, 1]$ on each edge (weight 0 \equiv no edge).

Definition (F., '18)

Fix $\mathbf{C} = (C_r)_{r \geq 1}$. A weighted graph \tilde{L} with vertex set A is a **C-weighted dependency graph** for the family $\{Y_\alpha, \alpha \in A\}$ if, for any $\alpha_1, \dots, \alpha_r$ in A ,

$$|\kappa(Y_{\alpha_1}, \dots, Y_{\alpha_r})| \leq C_r \mathcal{M}(\tilde{L}[\alpha_1, \dots, \alpha_r]).$$

Intuition: the smaller the edge weights are, the smaller the cumulant should be. The **edge weights quantify the dependencies** between variables.

⚠ Unlike for usual dependency graphs, **proving that something is a weighted dependency graph needs work!**

⚠ This is a **simplified version** of the definition; some of the applications need a more general but more technical version.

A normality criterion for weighted dependency graphs

Setting: for each n ,

- $\{Y_{n,i}, 1 \leq i \leq N_n\}$ is a family of bounded random variables; $|Y_{n,i}| < M$ a.s.
- we have a \mathbf{C} -weighted dependency graph \tilde{L}_n with weighted maximal degree $D_n - 1$ (with a sequence $\mathbf{C} = (C_r)_{r \geq 1}$ independent of n).
- we set $X_n = \sum_{i=1}^{N_n} Y_{n,i}$ and $\sigma_n^2 = \text{Var}(X_n)$.

A normality criterion for weighted dependency graphs

Setting: for each n ,

- $\{Y_{n,i}, 1 \leq i \leq N_n\}$ is a family of bounded random variables; $|Y_{n,i}| < M$ a.s.
- we have a \mathbf{C} -weighted dependency graph \tilde{L}_n with weighted maximal degree $D_n - 1$ (with a sequence $\mathbf{C} = (C_r)_{r \geq 1}$ independent of n).
- we set $X_n = \sum_{i=1}^{N_n} Y_{n,i}$ and $\sigma_n^2 = \text{Var}(X_n)$.

Theorem (F., '18)

Assume that $\left(\frac{N_n}{D_n}\right)^{1/s} \frac{D_n}{\sigma_n} \rightarrow 0$ for some integer s . Then X_n is asymptotically normal.

A normality criterion for weighted dependency graphs

Setting: for each n ,

- $\{Y_{n,i}, 1 \leq i \leq N_n\}$ is a family of bounded random variables; $|Y_{n,i}| < M$ a.s.
- we have a \mathbf{C} -weighted dependency graph \tilde{L}_n with weighted maximal degree $D_n - 1$ (with a sequence $\mathbf{C} = (C_r)_{r \geq 1}$ independent of n).
- we set $X_n = \sum_{i=1}^{N_n} Y_{n,i}$ and $\sigma_n^2 = \text{Var}(X_n)$.

Theorem (F., '18)

Assume that $\left(\frac{N_n}{D_n}\right)^{1/s} \frac{D_n}{\sigma_n} \rightarrow 0$ for some integer s . Then X_n is asymptotically normal.

Stability by powers

Setting:

- Let $\{Y_\alpha, \alpha \in A\}$ be r.v. with \mathbf{C} -weighted dependency graph \tilde{L} ;
- fix an integer $m \geq 2$;
- for a multiset $B = \{\alpha_1, \dots, \alpha_m\}$ of elements of A , denote

$$Y_B := Y_{\alpha_1} \cdots Y_{\alpha_m}.$$

Stability by powers

Setting:

- Let $\{Y_\alpha, \alpha \in A\}$ be r.v. with \mathbf{C} -weighted dependency graph \tilde{L} ;
- fix an integer $m \geq 2$;
- for a multiset $B = \{\alpha_1, \dots, \alpha_m\}$ of elements of A , denote

$$\mathbf{Y}_B := Y_{\alpha_1} \cdots Y_{\alpha_m}.$$

Proposition

The set of r.v. $\{\mathbf{Y}_B\}$ has a $\mathbf{C}^{(m)}$ -weighted dependency graph \tilde{L}^m , where

$$\text{wt}_{\tilde{L}^m}(\mathbf{Y}_B, \mathbf{Y}_{B'}) = \max_{\alpha \in B, \alpha' \in B'} \text{wt}_{\tilde{L}}(Y_\alpha, Y_{\alpha'}),$$

where $\mathbf{C}^{(m)}$ depends only on \mathbf{C} and m .

Convention: $\text{wt}_{\tilde{L}}(Y_\alpha, Y_\alpha) = 1$.

Stability by powers

Setting:

- Let $\{Y_\alpha, \alpha \in A\}$ be r.v. with \mathbf{C} -weighted dependency graph \tilde{L} ;
- fix an integer $m \geq 2$;
- for a multiset $B = \{\alpha_1, \dots, \alpha_m\}$ of elements of A , denote

$$\mathbf{Y}_B := Y_{\alpha_1} \cdots Y_{\alpha_m}.$$

Proposition

The set of r.v. $\{\mathbf{Y}_B\}$ has a $\mathbf{C}^{(m)}$ -weighted dependency graph \tilde{L}^m , where

$$\text{wt}_{\tilde{L}^m}(\mathbf{Y}_B, \mathbf{Y}_{B'}) = \max_{\alpha \in B, \alpha' \in B'} \text{wt}_{\tilde{L}}(Y_\alpha, Y_{\alpha'}),$$

where $\mathbf{C}^{(m)}$ depends only on \mathbf{C} and m .

This helps in proving correctness of weighted dependency graph!

Transition

- 1 Dependency graphs
 - A motivating example: substrings in random words
 - An asymptotic normality criterion
- 2 Weighted dependency graphs
 - Definition and an extended normality criterion
 - Back to subwords: Markovian texts

A weighted dependency graph for Markov chain

Setting:

- Let $(w_i)_{i \geq 1}$ be an irreducible aperiodic **Markov chain** on a finite space state \mathcal{A} ;
- Assume w_1 is distributed with the stationary distribution π ;
- Set $Z_{i,s} = \mathbf{1}_{w_i=s}$.

A weighted dependency graph for Markov chain

Setting:

- Let $(w_i)_{i \geq 1}$ be an irreducible aperiodic **Markov chain** on a finite space state \mathcal{A} ;
- Assume w_1 is distributed with the stationary distribution π ;
- Set $Z_{i,s} = \mathbf{1}_{w_i=s}$.

Proposition

We have a **weighted dependency graph** \tilde{L} with $\text{wt}_{\tilde{L}}(\{Z_{i,s}, Z_{j,t}\}) = |\lambda_2|^{j-i}$ (for $i < j$), where λ_2 is the second eigenvalue of the transition matrix.

Concretely, this means that, for $i_1 < \dots < i_r$,

$$|\kappa(Z_{i_1, s_1}, \dots, Z_{i_r, s_r})| \leq C_r \lambda_2^{i_r - i_1}.$$

It turns out that this was proved by Saulis and Statulevičius ('90)!

A weighted dependency graph for Markov chain

Setting:

- Let $(w_i)_{i \geq 1}$ be an irreducible aperiodic **Markov chain** on a finite space state \mathcal{A} ;
- Assume w_1 is distributed with the stationary distribution π ;
- Set $Z_{i,s} = \mathbf{1}_{w_i=s}$.

Proposition

We have a **weighted dependency graph** \tilde{L} with $\text{wt}_{\tilde{L}}(\{Z_{i,s}, Z_{j,t}\}) = |\lambda_2|^{j-i}$ (for $i < j$), where λ_2 is the second eigenvalue of the transition matrix.

Corollary (using the stability by product)

We have a **weighted dependency graph** \tilde{L}^m for monomials $Z_{I;S} := Z_{i_1,s_1} \cdots Z_{i_m,s_m}$, with $\text{wt}_{\tilde{L}^m}(Z_{I;S}, Z_{J;T}) = |\lambda_2|^{\text{md}(I,J)}$, where $\text{md}(I,J)$ is the minimal distance between I and J .

Subword occurrences in Markovian text (1/2)

Let $(w_i)_{i \geq 1}$ be a Markov chain as before and fix a pattern (= a word) u of length ℓ on \mathcal{A} .

For $I = \{i_1, \dots, i_\ell\} \subset \mathbb{N}$ ($i_1 < \dots < i_\ell$), we set

$$\begin{aligned} Y_I &= \mathbf{1}[u \text{ occurs at position } I \text{ in } \mathbf{w}]; \\ &= Z_{i_1, u_1} \cdots Z_{i_\ell, u_\ell}. \end{aligned}$$

Subword occurrences in Markovian text (1/2)

Let $(w_i)_{i \geq 1}$ be a Markov chain as before and fix a pattern (= a word) u of length ℓ on \mathcal{A} .

For $I = \{i_1, \dots, i_\ell\} \subset \mathbb{N}$ ($i_1 < \dots < i_\ell$), we set

$$\begin{aligned} Y_I &= \mathbf{1}[u \text{ occurs at position } I \text{ in } \mathbf{w}]; \\ &= Z_{i_1, u_1} \cdots Z_{i_\ell, u_\ell}. \end{aligned}$$

We have a **weighted dependency graph** for $(Y_I, I \in \binom{[n]}{\ell})$, which is a restriction of the one for the $Z_{I,S}$.

Subword occurrences in Markovian text (2/2)

Let $X_n = \sum_I Y_I$ be the number of occurrences of u in a Markovian text \mathbf{w} . Recall that $(Y_I, I \in \binom{[n]}{\ell})$ admits a weighted dependency graph.

Can we apply the normality criterion?

Subword occurrences in Markovian text (2/2)

Let $X_n = \sum_I Y_I$ be the **number of occurrences of u in a Markovian text w** . Recall that $(Y_I, I \in \binom{[n]}{\ell})$ admits a weighted dependency graph.

Can we apply the **normality criterion**? $M = 1$, $N_n = \binom{n}{\ell}$, and...

degree Fix $I = \{i_1, \dots, i_\ell\}$, we have

$$\sum_J \lambda_2^{\text{md}(I,J)} \leq \sum_{t,s} \left(\sum_J \lambda_2^{|i_t - j_s|} \right) \leq \sum_{t,s} \binom{n}{\ell-1} \sum_j \lambda_2^{|i_t - j|} = \mathcal{O}(n^{\ell-1}).$$

The maximal weighted degree D_n is $\mathcal{O}(n^{\ell-1})$.

variance $\sigma_n = \sqrt{\text{Var}(X_n)} = (C + o(1))n^{\ell-1/2}$, for a computable constant C (Bourdon, Vallée, '01).

Subword occurrences in Markovian text (2/2)

Let $X_n = \sum_I Y_I$ be the number of occurrences of u in a Markovian text w . Recall that $(Y_I, I \in \binom{[n]}{\ell})$ admits a weighted dependency graph.

Can we apply the normality criterion? $M = 1$, $N_n = \binom{n}{\ell}$, and...

degree Fix $I = \{i_1, \dots, i_\ell\}$, we have

$$\sum_J \lambda_2^{\text{md}(I,J)} \leq \sum_{t,s} \left(\sum_J \lambda_2^{|i_t - j_s|} \right) \leq \sum_{t,s} \binom{n}{\ell-1} \sum_j \lambda_2^{|i_t - j|} = \mathcal{O}(n^{\ell-1}).$$

The maximal weighted degree D_n is $\mathcal{O}(n^{\ell-1})$.

variance $\sigma_n = \sqrt{\text{Var}(X_n)} = (C + o(1))n^{\ell-1/2}$, for a computable constant C (Bourdon, Vallée, '01).

→ when $C > 0$, the normality criterion satisfied for $s = 3$.

Conclusion: when $C > 0$, the number X_n of occurrences of u in a Markovian text w is asymptotically normal.

(Answers partially a question of Bourdon–Vallée, '01).

Conclusion

- **Dependency graphs** are a powerful simple **tool to prove asymptotic normality**, particularly for substructure counts in models exhibiting some **independence**;
- We proposed an extension to handle models **without independence, but with weak dependencies**.
- **Plenty of applications** (both for the initial framework and for the extended one)!

Conclusion

- **Dependency graphs** are a powerful simple **tool to prove asymptotic normality**, particularly for substructure counts in models exhibiting some **independence**;
- We proposed an extension to handle models **without independence, but with weak dependencies**.
- **Plenty of applications** (both for the initial framework and for the extended one)!

Thank you for your attention!