### $\mathsf{Mod}\text{-}\phi$ convergence I

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# Central limit theorem (CLT) and beyond

- Standard CLT: renormalized sum of i.i.d. variables with finite variance tends towards a Gaussian distribution.
- Many relaxation of the i.i.d. hypothesis: CLT for Markov chains, martingales, mixing processes, m-dependent sequence, "associated" random variables...

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- We often have companion theorems: deviation probability, concentration inequalities, local limit theorem, speed of convergence. . .

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- Standard CLT: renormalized sum of i.i.d. variables with finite variance tends towards a Gaussian distribution.
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- We often have companion theorems: deviation probability, concentration inequalities, local limit theorem, speed of convergence...

But the companion theorems need extra effort to prove. Philosophy: Mod- $\phi$  is a universality class beyond the CLT, which implies some companion theorems.

## $\mathsf{Mod}\text{-}\phi$ convergence: definition

### Setting:

- D a domain of  $\mathbb{C}$  containing 0.
- ullet  $\phi$  infinite divisible distribution with Laplace transform  $\exp(\eta(z))$  on D.

### Definition (Nikeghbali, Kowalski)

A sequence of real r.v.  $(X_n)$  converges mod- $\phi$  on D with parameter  $t_n \to \infty$  and limiting function  $\psi$  if, locally uniformly on D,

$$\exp(-t_n \eta(z)) \mathbb{E}(e^{zX_n}) \to \psi(z),$$
 (1)

#### Informal interpretation:

- $X_n = t_n$  independent copies of  $\phi$  + perturbation encoded in  $\psi$ .
- instead of renormalizing the variables as in CLT, we renormalized the Fourier/Laplace transform to get access to the next term.

(this notion has some similarity with Hwang's quasi-powers.)

# $\mathsf{Mod} ext{-}\phi$ convergence implies a CLT

### Proposition

If  $(X_n)$  converges mod- $\phi$  on D with parameter  $t_n$ , then

$$Y_n = \frac{X_n - t_n \eta'(0)}{\sqrt{t_n \eta''(0)}} \longrightarrow_d \mathcal{N}(0,1).$$

Proof: easy, use the mod- $\phi$  estimate to show that  $\mathbb{E}(e^{\zeta Y_n})$  converges pointwise to  $e^{\zeta^2/2}$ .

Philosophy: Many classical ways of proving CLTs can be adapted to prove  $\text{mod-}\phi$  convergence.

(In particular, in all examples in the next few slides, the CLT is a well-known result.)

# Outline of today's talk

- $oldsymbol{1}$  Introduction: CLT and mod- $\phi$  convergence
- $oxed{2}$  Examples of mod- $\phi$  convergence sequences
  - ullet How to prove mod- $\phi$  convergence
- Companion theorems
  - Speed of convergence
  - Deviation and normality zone

# Examples with an explicit generating function (1/3)

We start with a trivial example.

Let  $Y_1, Y_2,...$  be i.i.d. with law  $\phi$  and  $W_n$  a sequence of r.v., independent from the Y, whose Laplace transform converges to that of W on D. Set  $X_n = W_n + \sum_{i=1}^n Y_i$ . Then

$$\mathbb{E}(\mathrm{e}^{zX_n}) = \mathrm{e}^{n\,\eta(z)}\,\mathbb{E}(\mathrm{e}^{zW_n}) = \mathrm{e}^{n\,\eta(z)}\,(\mathbb{E}(\mathrm{e}^{zW}) + o(1)).$$

Thus  $X_n$  converges mod- $\phi$  with parameters  $t_n = n$  and limiting function  $\psi(z) = \mathbb{E}(e^{zW})$ .

# Examples with an explicit generating function (2/3)

Let  $X_n$  be the number of cycles in a uniform random permutation.

$$\mathbb{E}[\mathrm{e}^{zX_n}] = \prod_{i=1}^n \left(1 + \frac{\mathrm{e}^z - 1}{i}\right) = \mathrm{e}^{H_n(\mathrm{e}^z - 1)} \, \prod_{i=1}^n \frac{1 + \frac{\mathrm{e}^z - 1}{i}}{\mathrm{e}^{\frac{\mathrm{e}^z - 1}{i}}}.$$

where  $H_n = \sum_{i=1}^n \frac{1}{i} = \log n + \gamma + \mathcal{O}(n^{-1})$ .

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where  $H_n = \sum_{i=1}^n \frac{1}{i} = \log n + \gamma + \mathcal{O}(n^{-1})$ . The product on the right-hand side converges locally uniformly on  $\mathbb{C}$  to an infinite product, which turns out to be related to the  $\Gamma$  function,

$$\mathbb{E}[\mathrm{e}^{zX_n}]\,\mathrm{e}^{-(\mathrm{e}^z-1)\log n}\to\mathrm{e}^{\gamma\,(\mathrm{e}^z-1)}\prod_{i=1}^{\infty}\frac{1+\frac{\mathrm{e}^z-1}{i}}{\mathrm{e}^{\frac{\mathrm{e}^z-1}{i}}}=\frac{1}{\Gamma(\mathrm{e}^z)}$$

locally uniformly, *i.e.*, one has mod-Poisson convergence on  $\mathbb{C}$  with parameters  $t_n = \log n$  and limiting function  $1/\Gamma(e^z)$ .

# Examples with an explicit generating function (3/3)

Other examples with explicit generating functions:

•  $\log(|\det(\operatorname{Id}-U_n)|)$  where  $U_n$  is an unitary Haar-distributed random matrices. It converges mod-Gaussian on  $\{\operatorname{Re}(z)>-1\}$  with parameter  $\frac{\log n}{2}$  and limiting function  $\Psi_1(z)=\frac{G(1+z/2)^2}{G(1+z)}$  (G is the G-Barnes function).

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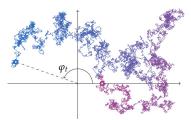
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- $\approx$  Let  $M_n$  be a GUE matrix. Then  $|\det(M_n)| \mathbb{E}(|\det(M_n)|)$  converges mod-Gaussian on  $\{|z| < 1\}$  with parameter  $t_n \sim \frac{1}{2}\log(n)$  and same limiting function  $\Psi_1(z)$  (Döring, Eichelsbacher, 2013).

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- $\varphi_t$  is the winding number of a Brownian motion starting at 1. It "converges mod-Cauchy on  $i\mathbb{R}$ " with parameter  $\frac{\log(8t)}{2}$  and limiting function  $\Psi_2(i\zeta) = \frac{\sqrt{\pi}}{\Gamma((|\zeta|+1)/2)}$ .



# Examples with an explicit **bivariate** generating function (overview)

Number  $\omega(k)$  of prime divisors of the integer k

$$\sum_{k \geq 1} \frac{\mathrm{e}^{z\omega(k)}}{k^s} = \prod_{p} \left( 1 + \frac{\mathrm{e}^z}{p^s(1 - p^{-s})} \right).$$

 $\Omega_n = \omega(k)$ , for a uniform random positive integer  $k \leq n$ .

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 $\Omega_n = \omega(k)$ , for a uniform random positive integer  $k \leq n$ .

Number of ascents  $A_n$  in a random permutation of size n

$$\sum_{n>1} \mathbb{E}(e^{zA_n})t^n = \frac{e^z - 1}{e^z - e^{t(e^z - 1)}}.$$

# Examples with an explicit **bivariate** generating function (overview)

Number  $\omega(k)$  of prime divisors of the integer k

$$\sum_{k \geq 1} \frac{\mathrm{e}^{z\omega(k)}}{k^s} = \prod_p \left( 1 + \frac{\mathrm{e}^z}{p^s(1-p^{-s})} \right). \quad \Omega_n \text{ converges mod-Poisson}$$

 $\Omega_n = \omega(k)$ , for a uniform random positive integer  $k \leq n$ .

Number of ascents  $A_n$  in a random permutation of size n

$$\sum_{n\geq 1}\mathbb{E}(\mathrm{e}^{zA_n})t^n=\frac{\mathrm{e}^z-1}{\mathrm{e}^z-\mathrm{e}^{t(e^z-1)}}.\quad A_n\text{ "converges mod-}U([0,1])\text{"}$$

In both cases one can extract the Laplace transform of  $\Omega_n$  or  $A_n$  by a path integral and study asymptotics.

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## A central limit theorem due to Harper

### Theorem (Harper, 1967)

Let  $X_n$  be a  $\mathbb{N}$ -valued random variable such that  $P_n(t) = \mathbb{E}(t^{X_n})$  has nonpositive real roots. Assume  $\text{Var}(X_n) \to \infty$ . Then

$$\frac{X_n - \mathbb{E}(X_n)}{\sqrt{\mathsf{Var}(X_n)}} \longrightarrow_d \mathcal{N}(0,1).$$

Example:  $X_n$  is the number of blocks of a uniform random set-partitions. (One can prove that

$$P_n(t)e^t = \operatorname{cst}_n t \frac{d}{dt} (P_{n-1}(t)e^t)$$

and apply Rolle's theorem inductively.)

# Mod-Gaussian convergence in Harper's theorem

### Theorem (FMN, 2013-2017)

Let  $X_n$  be a  $\mathbb{N}$ -valued random variable such that  $P_n(t) = \mathbb{E}(t^{X_n})$  is a polynomial with nonpositive real roots. Denote  $\sigma_n^2 = \operatorname{Var}(X_n)$  and  $L_n^3 = \kappa_3(X_n)$  the second and third cumulants of  $X_n$  and assume  $1 \ll L_n \ll \sigma_n \ll L_n^2$ .

Then  $\frac{X_n - \mathbb{E}(X_n)}{L_n}$  converges mod-Gaussian on  $\mathbb{C}$  with parameters  $t_n = \frac{\sigma_n^2}{L_n^2}$  and limiting function  $\psi = \exp(z^3/6)$ .

Idea of proof:  $X_n$  write as a sum of  $N_n$  Bernoulli variables  $B_k$  (of unknown parameters). Thus

$$\mathbb{E}(\mathrm{e}^{zX_n}) = \prod_{k=1}^{N_n} \mathbb{E}(\mathrm{e}^{zB_k})$$

and we do Taylor expansions on the right-hand side.

### Mod-Gaussian convergence in Harper's theorem

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Example:  $X_n$  is the number of blocks of a uniform random set-partitions. (The third cumulant estimate is not trivial.)

# Adapting the method of moments (1/3)

Instead of moments we use cumulants  $\kappa_r(X_n)$ . If X is a random variable, its cumulants are the coefficients of

$$\log \mathbb{E}[e^{zX}] = \sum_{r=1}^{\infty} \frac{\kappa^{(r)}(X)}{r!} z^{r}.$$

First cumulants:

$$egin{aligned} \kappa_1(X) &:= \mathbb{E}(X), \ \kappa_2(X) &:= \mathsf{Var}(X,Y) = \mathbb{E}(X^2) - \mathbb{E}(X)^2 \ \kappa_3(X) &:= \mathbb{E}(X^3) - 3\mathbb{E}(X^2)\mathbb{E}(X) + 2\mathbb{E}(X)^3. \end{aligned}$$

Fact:  $Y_n$  converge in distribution to  $\mathcal{N}(0,1)$  if  $Var(Y_n) \to 1$  and all other cumulants tend to 0.

# Adapting the method of moments (2/3)

### Definition (uniform control on cumulants)

A sequence  $(S_n)$  admits a uniform control on cumulants with DNA  $(D_n, N_n, A)$  and limits  $\sigma^2$  and L if  $D_n = o(N_n)$ ,  $N_n \to +\infty$  and

$$\forall r \geq 2, \ |\kappa^{(r)}(S_n)| \leq N_n (2D_n)^{r-1} r^{r-2} A^r; \frac{\kappa^{(2)}(S_n)}{N_n D_n} = (\sigma_n)^2 \to_{n \to \infty} \sigma^2; \qquad \frac{\kappa^{(3)}(S_n)}{N_n (D_n)^2} = L_n \to_{n \to \infty} L.$$

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### Proposition

Take  $(S_n)$  admits a uniform control on cumulants with  $\sigma^2 > 0$ . Then,  $X_n := \frac{S_n - \mathbb{E}[S_n]}{(N_n)^{\frac{1}{3}}(D_n)^{\frac{2}{3}}}$  converges mod-Gaussian on  $\mathbb{C}$ , with  $t_n = (\sigma_n)^2 \left(\frac{N_n}{D_n}\right)^{\frac{1}{3}}$  and limiting function  $\psi(z) = \exp(\frac{Lz^3}{6})$ .

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#### Remark

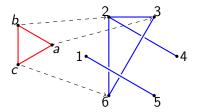
Uniform bounds on cumulants have been studied (in more generality) by Saulis and Statulevičius (1991) (see Döring-Eichelsbacher 2012, 2013, for numerous applications).

In this context, we don't have new theoretical results, but new examples.

# Adapting the method of moments (3/3): a new example

If  $F=(V_F,E_F)$  and  $G=(V_G,E_G)$  are finite graphs, a copy of F in G is a map  $\psi:V_F\to V_G$  such that

$$\forall e = \{x, y\} \in E_F, \ \{\psi(x), \psi(y)\} \in E_G.$$



#### Proposition

The number of copies of a fixed F in G(n,p) (p fixed) admits a uniform control on cumulants with DNA  $(n^{|V_G|-2}, n^{|V_G|}, 1)$  and  $\sigma^2 > 0$ .

(behind this: dependency graphs, more on that and more examples tomorrow!)

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### **Transition**

Reminder: if  $X_n$  converges mod- $\phi$ , then  $Y_n = \frac{X_n - t_n \eta'(0)}{\sqrt{t_n \eta''(0)}}$  converges to a standard Gaussian, i.e., for a fixed y,

$$\lim_{n\to\infty} \mathbb{P}(Y_n \ge y) = \frac{1}{\sqrt{2\pi}} \int_y^\infty e^{-u^2/2} du =: F_{\mathcal{N}}(y).$$
 (CLT)

### Main questions

Speed of convergence What is the error term (uniformly in y) in (CLT)? Deviation probability What if  $y \to \infty$ ? The limit is 0 but can we give an equivalent?

### A first bound for the speed of convergence

### Proposition (FMN, 2013-2017)

Let  $X_n$  converges mod- $\phi$  on a domain D containing  $i\mathbb{R}$ . Assume  $\phi$  non-lattice. Then

$$\mathbb{P}(Y_n \geq y) = F_{\mathcal{N}}(y) + \frac{\psi'(0)}{\sqrt{t_n \eta''(0)}} F_1(y) + \frac{\eta'''(0)}{6\sqrt{t_n (\eta''(0))^3}} F_2(y) + o(\frac{1}{\sqrt{n}}),$$

for explicit functions  $F_1(y)$  and  $F_2(y)$  (Gaussian integrals). In particular, the error term in (CLT) is  $\mathcal{O}(t_n^{-1/2})$  and we have an equivalent unless  $\psi'(0) = \eta'''(0) = 0$ .

 $\rightarrow$  Tight bounds for log(det(Id  $-U_n$ )), for  $A_n$ , but not for triangle count (see later). . .

# Bound on speed of convergence: ideas of proof

(Close to Feller, 1971, for the i.i.d. case.)

Standard tool in this context: Berry's inequality for centered variables

$$|F(y)-G(y)|\leq \frac{1}{\pi}\int_{-T}^{T}\left|\frac{f^*(\zeta)-g^*(\zeta)}{\zeta}\right|\,d\zeta+\frac{24m}{\pi T}.$$

F and G are distribution functions;  $f^*$  and  $g^*$  the Fourier transform of the corresponding laws; m a bound on the density g.

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Take  $F_n(y) = \mathbb{P}(Y_n \geq y)$  and

$$G_n(y) = \int_{-\infty}^{y} \left( 1 + \frac{\psi'(0)}{\sqrt{t_n \eta''(0)}} u + \frac{\eta'''(0)}{6\sqrt{t_n (\eta''(0))^3}} (u^3 - 3u) \right) g(u) du.$$

The mod- $\phi$  estimate allows you to control the integral for  $T = \Delta t_n^{1/2}$ .

(For  $\zeta \ll t_n^{1/2}$ ,  $f^*(\zeta) \sim g^*(\zeta)$ , for  $\zeta \approx t_n^{1/2}$ , both terms are small.)

Make n tends to infinity and then  $\Delta$ .

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# Speed of convergence for triangles in random graphs

Let  $T_n$  be the number of copies of  $F = K_3$  in G(n, p).

- Our bound gives an error term  $\mathcal{O}(n^{-1/3})$ .
- With a result of Rinott (1994), we can get  $O(n^{-1})$  (see also Krokowski, Reichenbachs and Thaele, 2015).

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#### Question

Can we improve our bounds in this case?

### A better bound from uniform control on cumulants

Proposition (Saulis, Statelivičius, 1991, FMN, 2017)

Let  $(S_n)$  be a sequence with a uniform control on cumulants with DNA  $(D_n, N_n, A)$  with  $\sigma^2 > 0$ .

(In particular,  $|\kappa^{(r)}(S_n)| \leq N_n (2D_n)^{r-1} r^{r-2} A^r$ .)

Then the error term in (CLT) is  $\mathcal{O}(t_n^{-3/2}) = \mathcal{O}(\sqrt{D_n/N_n})$ .

In case of triangles, we get  $\mathcal{O}(n^{-1})$  as Rinott (1994) or Krokowski, Reichenbachs and Thaele (2015).

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Proof: again Berry's inequality

$$|F(x)-G(x)|\leq \frac{1}{\pi}\int_{-T}^{T}\left|\frac{f^*(\zeta)-g^*(\zeta)}{\zeta}\right|\ d\zeta+\frac{24m}{\pi T}.$$

but we have a better control on  $f^*(\zeta)$  and thus we can choose  $T=t_n^{3/2}$ .

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Our statement is a bit more general: holds in the context of mod-stable convergence with additional control of the Laplace transform.

Example: winding number  $\varphi_t$  of a Brownian motion converges to a Cauchy law after renormalization at speed  $\mathcal{O}(t_n) = \mathcal{O}(\log n)$ .

### Deviation probability

### Theorem (FMN, 2013-2017)

Assume  $X_n$  converges mod- $\phi$  ( $\phi$  non-lattice) on a strip  $\{|Re(z)| \leq C\}$ . Let  $x_n$  bounded by C with  $x_n \gg t_n^{-1/2}$ . Then

$$\mathbb{P}\big(X_n - t_n \eta'(0) \geq t_n x_n\big) \sim_{n \to \infty} \frac{\exp(-t_n F(x_n))}{h_n \sqrt{2\pi t_n \eta''(h_n)}} \psi(h_n)(1 + o(1)).$$

Here  $F(x) = \sup_{h \in \mathbb{R}} (hx - \eta(h))$  is the Legendre Fenchel transform of  $\eta$  and  $h_n$  is the maximizer for  $F(x_n)$ .

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$$\mathbb{P}\big(X_n-t_n\eta'(0)\geq t_nx_n\big)\sim_{n\to\infty} \frac{\exp(-t_nF(x_n))}{h_n\sqrt{2\pi t_n\eta''(h_n)}}\psi(h_n)(1+o(1)).$$

Standard proof strategy: applying speed of convergence result to the exponentially tilted variables  $\widetilde{X}_n$ :

$$\mathbb{P}[\widetilde{X}_n \in du] = \frac{\mathbb{E}^{hu}}{\varphi_{X_n}(h)} \mathbb{P}[X_n \in du].$$

 $\widetilde{X}_n$  also converge mod- $\phi$ : its Laplace transform is simply  $\mathbb{E}[\mathrm{e}^{z\,\widetilde{X}_n}] = \frac{\mathbb{E}[\mathrm{e}^{(z+h)\,X_n}]}{\mathbb{E}[\mathrm{e}^{h\,X_n}]}.$ 

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Similar result for lattice distributions  $\phi$ : replace h in denominator by  $e^h - 1$ .

## Normality zone

#### Definition

We say that  $Y_n$  has a normality zone  $o(a_n)$  if (CLT) gives an equivalent of the tail probability for  $y = o(a_n)$  but not for  $y = O(a_n)$ .

### Proposition

Let  $X_n$  converges mod- $\phi$  on a strip  $\{|Re(z)| \leq C\}$ . Then the normality zone of  $\frac{X_n - t_n \eta'(0)}{\sqrt{t_n \eta''(0)}}$  is  $o(t_n^{1/2 - 1/m})$ , where  $m \geq 3$  is minimal such that  $\eta^{(m)} \neq 0$ . If  $\phi$  is Gaussian,  $m = \infty$  by convention, but we need to assume that  $\psi \not\equiv 1$ .

• Let  $T_n$  be the number of copies of  $F = K_3$  in G(n, p). Then

$$\mathbb{P}\big[T_n \ge n^3 p^3 + n^2 (v - 3p^3)\big] \sim \sqrt{\frac{9p^5 (1-p)}{\pi v^2}} \exp\left(-\frac{v^2}{36 p^5 (1-p)} + \frac{(7-8p) v^3}{324 n p^8 (1-p)^2}\right)$$
 for  $1 \ll v = O(n^{2/3})$ .

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 for  $1 \ll v = O(n^{2/3})$ .

• Let  $A_n$  be the number of ascents in a random permutation of size n.

$$\mathbb{P}\left[A_n \ge \frac{n+1}{2} + \sqrt{\frac{n+1}{12}} y\right] = \frac{(1+o(1))}{y\sqrt{2\pi}} \exp\left(-\frac{y^2}{2} + \frac{y^4}{120(n+1)}\right)$$

for any positive y with  $y = o(n^{\frac{5}{12}})$ .

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- Let  $T_n$  be the number of copies of  $F = K_3$  in G(n, p). Then  $\mathbb{P} \left[ T_n \ge n^3 p^3 + n^2 (v 3p^3) \right] \sim \sqrt{\frac{9p^5 (1-p)}{\pi v^2}} \exp \left( -\frac{v^2}{36 p^5 (1-p)} + \frac{(7-8p) v^3}{324 n p^8 (1-p)^2} \right)$
- Let  $A_n$  be the number of ascents in a random permutation of size n.

$$\mathbb{P}\left[A_n \ge \frac{n+1}{2} + \sqrt{\frac{n+1}{12}}\,y\right] = \frac{(1+o(1))}{y\sqrt{2\pi}}\,\exp\left(-\frac{y^2}{2} + \frac{y^4}{120(n+1)}\right)$$

for any positive y with  $y = o(n^{\frac{5}{12}})$ .

• Let  $U_n$  be Haar distributed in the unitary group U(n), one has: for  $x_n \gg (\log n)^{-1/2}$  bounded,

$$\mathbb{P}_n\Big[\,|\det(\mathsf{Id} - U_n)| \geq n^{\frac{x_n}{2}}\Big] = \frac{G(1+\frac{x_n}{2})^2}{G(1+x_n)}\,\frac{1}{x_n\,n^{\frac{x_n^2}{4}}\,\sqrt{\pi\log n}}\,(1+o(1)).$$

• Let  $T_n$  be the number of copies of  $F = K_3$  in G(n, p). Then

$$\mathbb{P}\big[T_n \ge n^3 p^3 + n^2 (\nu - 3p^3)\big] \sim \sqrt{\frac{9p^5(1-p)}{\pi \nu^2}} \, \exp\!\left(-\frac{\nu^2}{36 \, p^5(1-p)} + \frac{(7-8p) \, \nu^3}{324 \, n \, p^8(1-p)^2}\right)$$
 for  $1 \ll \nu = O(n^{2/3})$ .

• Let  $A_n$  be the number of ascents in a random permutation of size n.

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(We also have estimates for negative deviations in all cases.)

V. Féray (UZH)

### Conclusion

#### Future work:

- Concentration estimates, local limit theorems. . .
- Prove mod- $\phi$  convergence in other contexts where the CLT is known: martingales, Stein exchangeable pairs, linear statistics of determinental processes, mixing processes . . .

Tomorrow: dependency graphs, variants and mod-Gaussian convergence.